Basel II and Developing Countries: Diversification and Portfolio Effects \(^1\)

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Introduction

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Our concerns on the potential impact of the proposed new Basel Capital Accord (Basel II) were first expressed following the release of the second Consultative Paper (CP2) in January 2001. However, since that time a number of modifications have been made to the proposals that go some way to addressing these original concerns.

The most recent paper of ours on this subject was published in the *Financial Regulator* in September 2002. This paper reiterated our concerns about the potential impact of the proposals on developing and emerging economies, assessed the likely impact of the modifications announced by that time, and highlighted remaining areas of concern. These were twofold:

1. Widespread adoption of the IRB approach by internationally active banks would lead to a significant increase (decrease) in capital requirements for loans to lower (higher) rated borrowers. To the extent that the pricing and availability of international bank loans is influenced by the capital requirements that relate to them, this would imply a sharp increase in the cost and/or a reduction in the quantity of international lending to developing and emerging economies. Given the current very low levels of such lending, this raises the possibility of the current situation becoming ‘institutionalised’, so that, even if global conditions improve, the potential of international bank lending to contribute towards the development of poorer countries would be significantly reduced.

2. The use of market-sensitive measures of risk – as envisaged in the IRB approaches – is inherently pro-cyclical. The fact that capital requirements will move in conjunction with the business cycle implies an amplification of that cycle as loans ‘migrate’ between bands as circumstances improve or deteriorate. The natural tendency of market practitioners – including bankers – to underestimate risks in booms and overestimate risks in recessions will thus be formalised, and legitimised, in regulation. Thus, in an upturn, the perception of generally reduced risks would result in lower capital requirements, further strengthening this perception of lower risk, but perhaps resulting in a longer ‘boom’ period and the build-up of greater levels of potentially systemic risk. Conversely, in a downturn or recession, higher capital requirements, as determined by the IRB approach, would reduce further incentives to lend, and – coupled with the difficulty of raising capital in a recession - create the possibility of a ‘credit crunch’ wherein even potentially profitable business propositions are unable to attract funding. The danger is that a downturn is turned into a recession, or an existing recession lengthened or deepened.

These concerns about the potentially damaging impact of Basel II were viewed in the context of a more general analysis. This argued that that the major problems facing developing countries in their attempts to access international finance for purposes of growth and development were a) the current low level of all types of flows (particularly, but not exclusively, bank lending) and b) the increasingly short-term and pro-cyclical nature of these flows. (Griffith-Jones, 2002) Given our view of this

http://www.ids.ac.uk/ids/global/finance/ifpubs.html
discouraging general environment, it remains of serious concern that the proposals for Basel II may exacerbate, rather than attempt to counter, these damaging trends.

This paper will present the results of empirical work that we have undertaken to address the first point detailed above. We suggested in our most recent paper on this subject that one reason why capital requirements under the new proposals could be inappropriately high for developing and emerging economies, is that the benefits of international diversification are not taken into account. We suggested that, if it could be demonstrated that the correlation between developed/developed country lending was higher than that between developed/developing, then a case could be made that an internationally diversified loan portfolio, with a range of developed and developing country borrowers, would have a lower level of risk – in terms of the overall portfolio – than one which focused primarily on developed country lending. If this is, in fact, the case, then it would be possible – and certainly desirable – for the Basel Committee to incorporate the benefits of international diversification into the new Accord.

This argument is similar to that used to support the recent modifications (November, 2001) resulting in the flattening of the IRB curve, with respect to corporate lending. In the original proposals for January 2001 it implicitly assumed that the average asset correlation was 0.2. However following empirical research initiated by the Committee (Lopez 2001) a modification to the IRB formula was proposed so that the correlation coefficient would decline from 0.2 to 0.1 as probability of default (PD) increased. In essence, the argument is that a higher PD for a corporate reduces correlation, as bankruptcy/default may be the result of any number of non-systemic factors that would not necessarily have any impact on the prospects for other corporates.

The argument that asset correlation is variable is self-evident. Furthermore, the suggestion that this variability impacts upon the level of risk in an overall portfolio, and should therefore be reflected in capital requirements would also seem to have force. Consequently, we have followed this approach in our own empirical work, which, as we shall detail below, provides strong support for a similar modification of the IRB formula with respect to internationally diversified lending.

It has long been argued that one of the major benefits of investing in developing and emerging economies is their relatively low correlation with mature markets. Therefore our first hypothesis can be stated as follows:

\textbf{H1} – The degree of correlation between the real and financial sectors of developed economies is greater than that which exists between developed and developing economies.

We have tested this hypothesis of differential correlations, first with specific regard to international bank lending and profitability and, secondly, in a more general but supportive sense. All of our results offer significant support for the validity of this position. This has provided the basis for a second hypothesis, which relates specifically to the ongoing work of the Basel Committee:

\textbf{H2} - An international loan portfolio which is diversified across the developed, emerging and developing regions enjoys a more efficient risk/return trade-off – and
therefore lower overall portfolio level risk as measured by unexpected losses - than one focused exclusively on developed markets

In order to test this more specific hypothesis we have simulated levels of unexpected loss for two portfolios: one with a loan portfolio that is evenly distributed across developed and developing regions; the second with a portfolio that is distributed across only the developed regions. The results of these simulations provide convincing support for the second of our hypotheses. Suggesting that the level of unexpected loss that a portfolio focused on purely developed country borrowers would face in an extreme event, would be about twenty-five percent higher than a portfolio diversified across developed and developing countries.

The fact that the tests we have performed, using a variety of variables, over a range of time periods, all provide strong evidence in support of our diversification hypothesis, seems, to us, compelling. This evidence is further strengthened by the results of our simulations of loan portfolios, which, by employing a similar methodology to that used by the most sophisticated banks, demonstrate the beneficial impacts of international diversification, as they would be viewed by the major banks. Taken together, this evidence suggests that, so as to not unfairly penalise emerging and developing economies, the Basel Committee should closely examine the practicalities of incorporating the benefits of international diversification into the forthcoming final consultative paper. It is hoped that the evidence presented below will demonstrate the validity of this view.

The rest of the paper is structured as follows. Section I details the sources of data and methodology used, section II presents the results of the econometric work, section III presents a simulation of two loan portfolios, section IV explores the implications of our results and concludes. Technical details on the statistical and simulation work are contained in the appendices.
I. Data and Sources

Countries analysed:

*Developing Countries:* Argentina, Brazil, Chile, Ecuador, Mexico, Panama, Peru, Venezuela, Philippines, Korea, Malaysia, Thailand, Indonesia, Bulgaria, Poland, Russia, Nigeria, South Africa

*Developed Countries:* U.S. Japan, Germany, Spain, France, U.K. Italy, Canada

*Others:* Singapore, Ireland, Greece, Portugal, Finland

Variables analysed:

Table 1.

<table>
<thead>
<tr>
<th>Grouping</th>
<th>Code</th>
<th>Description</th>
<th>Time Period</th>
<th>Freq</th>
<th>Source</th>
</tr>
</thead>
<tbody>
<tr>
<td>Financial Sector</td>
<td>ROA</td>
<td>Return on Assets (banks)</td>
<td>1988-2001</td>
<td>Annual</td>
<td><em>The Banker</em></td>
</tr>
<tr>
<td>Financial Sector</td>
<td>ROC</td>
<td>Return on tier one capital (banks)</td>
<td>1988-2001</td>
<td>Annual</td>
<td><em>The Banker</em></td>
</tr>
<tr>
<td>Financial Sector</td>
<td></td>
<td>Syndicated Loans Spreads</td>
<td>93-02</td>
<td>Monthly</td>
<td>BIS</td>
</tr>
<tr>
<td>Bonds</td>
<td>GBI</td>
<td>Global Bond Index</td>
<td>87-02</td>
<td>Daily</td>
<td>JP Morgan/Reuters</td>
</tr>
<tr>
<td>Bonds</td>
<td>EMBI</td>
<td>Emerging Market Bond Index</td>
<td>87-02</td>
<td>Daily</td>
<td>JP Morgan/Reuters</td>
</tr>
<tr>
<td>Bonds</td>
<td>EMBI+</td>
<td>Emerging Market Bond Index Plus.</td>
<td>87-02</td>
<td>Daily</td>
<td>JP Morgan/Reuters</td>
</tr>
<tr>
<td>Stocks</td>
<td>IFC G</td>
<td>S&amp;P International Finance Corporation (Global)</td>
<td>90-02</td>
<td>Daily</td>
<td>IFC/S&amp;P</td>
</tr>
<tr>
<td>Stocks</td>
<td>IFC I</td>
<td>S&amp;P International Finance Corporation (Investable)</td>
<td>90-02</td>
<td>Daily</td>
<td>IFC/S&amp;P</td>
</tr>
<tr>
<td>Stocks</td>
<td>COMP</td>
<td>Developed countries listed above: composite stock indexes</td>
<td>90-02</td>
<td>Daily</td>
<td>Reuters</td>
</tr>
<tr>
<td>Macro</td>
<td>GDP</td>
<td>GDP Growth Rate</td>
<td>85-00</td>
<td>Six-Monthly</td>
<td>IMF, World Bank (Author’s own calculations)</td>
</tr>
<tr>
<td>Macro</td>
<td>GDP HP</td>
<td>Hodrick-Prescott decomposition of GDP</td>
<td>50-98</td>
<td>Annual</td>
<td>National Data (Author’s own calculations)</td>
</tr>
<tr>
<td>Macro</td>
<td>STIR</td>
<td>Short term nominal interest rate</td>
<td>85-00</td>
<td>Six-Monthly</td>
<td>National data (BIS) or IMF, IFS</td>
</tr>
<tr>
<td>Macro</td>
<td>STIRR</td>
<td>Short term real interest rate</td>
<td>85-00</td>
<td>Six-Monthly</td>
<td>National data (BIS) or IMF, IFS</td>
</tr>
</tbody>
</table>

3 The GBI consists of regularly traded, fixed-rate, domestic government bonds. The countries covered have liquid government debt markets, which are freely accessible to foreign investors. GBI excludes: floating rate notes, perps, bonds with less than one year maturity, bonds targeted at the domestic markets for tax reasons and bonds with callable, puttable or convertible features.

4 Included in the EMBI are US dollar denominated Brady bonds, Eurobonds, traded loans and local debt market instruments issued by sovereign and quasi-sovereign entities.

5 EMBI+ is an extension of the EMBI. The index tracks all of the external currency denominated debt markets of the emerging markets.

6 IFC G (Global) is an emerging equity market index produced in conjunction with S&P. The index does not take into account restrictions on foreign ownership that limit the accessibility of certain markets and individual stocks.

7 IFC I (Investable) is adjusted to reflect restrictions on foreign investments in emerging markets. Consequently, it represents a more accurate picture of the actual universe available to investors.
II. Results

All the statistical significance tests we have undertaken provide strong support for our first hypothesis. Crucially for the validity of our results, cumulative distribution function (CDF) tests were undertaken in each instance. The purpose of the tests was to establish, for any given level of correlation, the probabilities that the developed/developed series and the developed/developing series would have a lower level of correlation. The results of two of these tests are shown in figures 1 and 2 (the remaining results are contained in Annex 1) as an example of the fact that, in every instance, the developed/developed correlation dominates that of the developed/developing correlation.

That is, for any level of correlation (x), the probability that the actual correlation between developed and developing indicators is lower than x, is higher than the probability that the correlation between developed and developed indicators is lower than x.

The results in Table 2 offer further support for the first of our hypotheses, in both a general and a specific sense. The specific, financial sector, results are presented first, followed by evidence from other, more general sources.

Table 2.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Time-Period</th>
<th>Frequency</th>
<th>Developed/Developed Mean Correlation Coefficient</th>
<th>Developed/Developing Mean Correlation Coefficient</th>
<th>Test Statistic (H0: Mx=My) Critical Value of 0.05% one-tailed test in parentheses</th>
</tr>
</thead>
<tbody>
<tr>
<td>Syndicated</td>
<td>1993-2002</td>
<td>Monthly</td>
<td>0.37</td>
<td>0.14</td>
<td>3.33 (3.29)</td>
</tr>
<tr>
<td>ROA</td>
<td>1988-2001</td>
<td>Annual</td>
<td>0.10</td>
<td>-0.08</td>
<td>4.40 (3.29)</td>
</tr>
<tr>
<td>ROC</td>
<td>1988-2001</td>
<td>Annual</td>
<td>0.14</td>
<td>-0.11</td>
<td>6.92 (3.29)</td>
</tr>
<tr>
<td>GDP</td>
<td>1985-2000</td>
<td>Annual</td>
<td>0.44</td>
<td>0.02</td>
<td>9.08 (3.29)</td>
</tr>
<tr>
<td>GDP HP</td>
<td>1950-1998</td>
<td>Annual</td>
<td>0.35</td>
<td>0.02</td>
<td>9.41 (3.29)</td>
</tr>
<tr>
<td>STIR</td>
<td>1985-2000</td>
<td>Six-monthly</td>
<td>0.72</td>
<td>0.23</td>
<td>11.09 (3.29)</td>
</tr>
<tr>
<td>STIRR</td>
<td>1985-2000</td>
<td>Six-monthly</td>
<td>0.66</td>
<td>0.22</td>
<td>10.93 (3.29)</td>
</tr>
<tr>
<td>GBI-EMBI</td>
<td>1991-2002</td>
<td>Daily</td>
<td>0.78</td>
<td>0.53</td>
<td>5.45 (3.29)</td>
</tr>
<tr>
<td>GBI-EMBI</td>
<td>1991-1997</td>
<td>Daily</td>
<td>0.90</td>
<td>0.74</td>
<td>4.64 (3.29)</td>
</tr>
<tr>
<td>GBI-EMBI</td>
<td>1998-2002</td>
<td>Daily</td>
<td>0.42</td>
<td>0.09</td>
<td>5.87 (3.29)</td>
</tr>
<tr>
<td>IFCI-COMP</td>
<td>1990-2000</td>
<td>Daily</td>
<td>0.58</td>
<td>-0.15</td>
<td>7.83 (3.29)</td>
</tr>
<tr>
<td>IFCG-COMP</td>
<td>1990-2000</td>
<td>Daily</td>
<td>0.58</td>
<td>-0.17</td>
<td>8.06 (3.29)</td>
</tr>
</tbody>
</table>
As can be seen from Table 2, all the results were tested to ensure statistical
significance. In all cases, the results were significant at the 99.5% confidence level
and the null hypothesis that the average mean correlations of the two series were
equal (H0: Mx=My) was clearly rejected.

Discussion

As is clear from table 1, a wide variety of financial, market and macro variables have
been employed in our tests. Whilst it might be suggested that each of the variables we
have used could be criticized as imperfect in some way, we would argue strongly that
the possibility of distortions in the data are likely to be cancelled out, as they are
unlikely to be the result of common causes. Consequently, the fact that every
statistical test that we have performed, regardless of variable, time-period or
frequency, has pointed in the same direction, and all are clearly statistically significant
on a variety of tests, offers robust and unequivocal support for our first hypothesis.

In the case of spreads on syndicated bank loans, and adopting the reasonable
assumption that they are indicative of the risk associated with the loans – and
therefore a proxy for probability of default – it is clear that risks, as measured in this
way, have had a greater tendency to rise and fall together within the developed
regions than has been the case for the developed and developing regions.
Consequently, this first result would appear to offer support to our hypothesis. That is,
over the sample period of 1993 to 2002 a bank with a loan portfolio that was well
diversified across the major developed and developing regions, would have enjoyed
diversification benefits at the portfolio level: the correlation between the risks
associated with loans to each of these regions would have been lower than was the
case for a bank with a loan portfolio which focused only on developed markets.

Similarly, the fact that the profitability of banks in developed markets are slightly
negatively correlated with those in developing markets, whilst the profitability of
banks within developed markets are slightly positively correlated, provides further
support for our hypothesis of the benefits of diversification. Although there may be
many factors affecting the level of profitability of a country’s domestic banking
system, it seems reasonable to assume that one of the more significant factors would
be the incidence of non-performing loans in the domestic economy. More generally,
the health and consequent profitability of the country’s domestic economy must
plausibly impact strongly upon the profitability of its banking sector. Thus, over the
sample period, a bank lending to both banks and corporates across a wide range of
developed and developing countries would have obtained diversification benefits, at
the portfolio level, relative to a bank with a loan portfolio concentrated solely on
developed markets.

The results from the macro variables, whilst more general, give some indication of the
extent to which developed economies have tended to move in step with each other to
a far greater extent than have developed and developing economies. If we plausibly
assume that the incidence of non-performing loans (NPL) in an economy is, at least
partially, inversely related to the rate of GDP growth, then banks with an
internationally diversified portfolio would be less likely to experience sharp increases
in the incidence of NPLs in these markets simultaneously. Conversely, a bank that
focused entirely on the – more highly correlated – mature markets would have a
greater chance of experiencing such an outcome. Similar implications can be drawn if we take movements in short-term interest rates as a proxy for the business cycle – rising rates indicating the close of an upturn and vice versa – these results provide further evidence in support of our argument. As with GDP growth, the fact that business cycles – and therefore movements in short-term interest rates – are more correlated between developed countries than between developed and developing countries, suggests that the incidence of NPLs and defaults are likely to be more correlated in the former than the latter.

For many market practitioners, movement in government bond prices and yields are seen as a strong indicator of both economic fundamentals and market views on the economic prospects of each country. The fact that developed country bond prices move in step to a far greater extent than do developed and developing country prices, suggests a closer correlation between both economic fundamentals in developed countries and market sentiment towards them. The evidence of lower correlation between developed and developing stock markets also supports this view. To the extent that a country’s stock market reflects economic fundamentals and investor sentiment towards the country, a lower correlation between developed and developing countries provides further evidence in support of our first hypothesis.

The evidence presented above clearly supports our hypothesis that a bank’s loan portfolio that is diversified internationally between developed and developing country borrowers would benefit in terms of lower overall portfolio risk relative to one that focused exclusively on lending to developed countries. In order to test this hypothesis in the specific context of a bank’s loan portfolio a simulation exercise has been undertaken to assess the potential unexpected loss resulting from a portfolio diversified within developed countries, and one diversified across developed and developing regions.

III. Simulated Loan portfolios

The testing of our second hypothesis involves the construction of two simulated loan portfolios, which enables us to assess the probable level of unexpected loss in each portfolio. Thus we can directly compare the simulated behaviour of a portfolio diversified across developed and developing regions, with one focused solely on developed markets.

The basic context for our approach and the results obtained are detailed below. Appendix 2 contains more information, as well as technical details of the construction of the simulated portfolios.

Context

The fact that the quality of the credit portfolio of any bank can change at any time in the future means that there is a need to make frequent calculations of the expected losses that a bank could suffer, under a variety of situations. Given the constant changes in portfolio quality, it is unlikely that the computed preventive reserves will be the same for different periods. The difference between preventive reserves computed at different periods, (due to changing credit quality), is the cause of the potential losses to the bank - those that could erode their capital in extreme situations.
These losses are called “Unexpected Losses”. Our second hypothesis, in effect, states that the levels of unexpected loss for a portfolio that is diversified across developed and developing markets will be lower than that for a portfolio that focuses exclusively on developed markets. This hypothesis is supported, in principle, by the results of our statistical work above, which demonstrated the lower level of correlation between developed/developing markets than that which exists between developed/developed markets.

**Simulation**

The approach we employ represents a modification of the well-known CreditMetrics approach, which has been widely used to simulate unexpected losses in portfolios. Following a similar approach, two simulated portfolios were constructed: one with an even distribution of loans across the major developed and developing regions; the other with the loan portfolio evenly distributed across the developed regions. We then programmed an algorithm that simulated 10,000 different ‘quality scenarios’ that might impact on these portfolios, and so produce migration of loans between credit quality bands. Each quality scenario shows a change in the market value of the assets of the creditors in the portfolio, and therefore the difference between the initial and final credit quality can be assessed. Once the credit portfolio quality scenarios have been simulated, it is possible to compute the losses/gains that come from the difference between initial and final credit qualities.

The losses/gains obtained from the simulation process are used to build a histogram, which summarises the loss distribution of the credit portfolio. From this distribution a ‘value at risk’ (VaR) is defined from which we obtain the amount of unexpected losses from the portfolio. The unexpected losses divided by the total amount of the portfolio represent the percentage that, with a given probability, (defined by the chosen percentile) could be lost in an extreme event.

**Results**

The results obtained from our simulations offer strong support for our second hypothesis. The results are as follows:

**Table 3. Comparison of non-industrially diversified portfolios**

<table>
<thead>
<tr>
<th>Percentile</th>
<th>Loss value</th>
<th>Unexpected loss (%)</th>
<th>Percentile</th>
<th>Loss value</th>
<th>Unexpected loss (%)</th>
<th>Percentage Difference</th>
</tr>
</thead>
<tbody>
<tr>
<td>99.8</td>
<td>22,595,312</td>
<td>19.21</td>
<td>99.8</td>
<td>27,869,349</td>
<td>23.69</td>
<td>+23.34</td>
</tr>
<tr>
<td>99.9</td>
<td>26,390,246</td>
<td>22.44</td>
<td>99.9</td>
<td>32,187,075</td>
<td>27.36</td>
<td>+21.96</td>
</tr>
</tbody>
</table>

8 Developing: Africa and the Middle East; Asia and Pacific; developing Europe; Latin America. Developed: EU (non-EMU); EMU; Other Industrial; offshore centres.

9 There are, of course, many problems and critics of the VaR approach to risk management. See Zigrand (2001) and Persaud (2001) for example. However, it is beyond the scope of this paper to assess these issues. For the purposes of this research, our simulation is designed to demonstrate – in broad terms – the relative difference in unexpected losses that would be likely to occur in each portfolio, in a similar fashion to that currently practiced by many major, internationally active banks.
As can be seen from table 3, the unexpected losses simulated for the portfolio focused on developed country borrowers are, on average, almost twenty-three percent higher than for the portfolio diversified across developed and developing countries.

Discussion

The simulated loan portfolios constructed offer clear evidence that the benefits of international diversification produce a more efficient risk/return trade-off for banks at the portfolio level. Given that capital requirements are intended to deal with unexpected loss, the fact that the level of unexpected loss in our simulation is lower for a diversified than for an undiversified portfolio, suggests that – in order to accurately reflect the actual risks that banks may face – Basel II should take account of this effect.

It is, of course, always possible to question the assumptions which underpin any simulation. We have attempted to ensure that our assumptions are as reasonable as possible. One aspect that we considered in detail was that the decision to assume no industrial diversification within countries might prevent the benefits of such diversification in developed countries – which generally have a greater range of industries than do developing countries – from being taken into account. We concluded, however, that the potential benefits of such diversification may have traditionally been overstated. This position is supported by recent empirical work undertaken by the BIS. Using data from 105 Italian banks, over the period 1993-1999, Acharya et al (2002) test empirically for evidence in support of the theoretical benefits of industrial, sectoral and geographical diversification. The results, though somewhat surprising, would seem to offer support for both the assumptions that underpin the loan portfolio simulation (i.e. no industrial diversification) and, crucially, the general findings or our empirical work.

From the combined results on bank loan return and risk, we conclude that increased industrial loan diversification results in an inefficient risk-return trade-off for the (Italian) banks in our sample, and sectoral diversification results in an inefficient risk-return trade-off for banks with relatively high levels of risk. Geographical diversification on the other hand does result in an improvement in the risk-return trade-off for banks with low or moderate levels of risk. (op. cit: 5)

However, in order to be certain that the simulation results have not been biased by this assumption, a second series of simulations was undertaken. In this instance, both geographical and industrial diversification was assumed. As can be seen in table 4, this modification – which brings the simulation closer to real practice - has the effect of halving the level of unexpected loss in the portfolios; thus they are now closer to the 8% figure often encountered in the real world, and which forms the basis of the Basel Committee’s stated capital requirements for the system as a whole.

Table 4. Comparison of two simulated industrially diversified portfolios

<table>
<thead>
<tr>
<th>1. Diversified developed/developing</th>
<th>2. Diversified developed</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total Exposure = 117,625,333.00</td>
<td>Total Exposure = 117,625,333.00</td>
</tr>
<tr>
<td>Percentile</td>
<td>Loss value</td>
</tr>
<tr>
<td>99.8</td>
<td>15,111,321</td>
</tr>
<tr>
<td>99.9</td>
<td>15,358,788</td>
</tr>
</tbody>
</table>

| Percentage Difference | 16.90 | 16.94 |

The difference between the simulated unexpected losses in the two portfolios has also been reduced by this modification, although less so. However, at almost seventeen percent, on average, the difference remains highly significant, and so offers further evidence of the robustness of our results.

Another issue that we have given consideration to is the fact that correlations are not constant over time. The danger, of course, is that correlations within emerging markets increase dramatically in crises, as contagion spreads the crisis from one country or region to another. In this instance, it is possible that a portfolio diversified across a range of emerging and developing regions, might be hit simultaneously in each of these areas. However, while this may well be the common perception of emerging market behaviour in crises, it may only apply to a limited number of cases, which require specific preconditions to be in place; preconditions, which at the current time – and indeed at most times - do not apply. Kaminsky, Reinhart and Vegh (2002) examine two hundred years of financial crises, in both developed and developing countries, for evidence of contagion. They conclude that ‘fast and furious’ contagion of the type described above, and often viewed as inherent in emerging markets may occur, but only under certain circumstances. Of the major emerging market crises since 1980, the Mexican default of 1982, the Mexican devaluation of 1994, the devaluation of the Thai baht in 1997 and the Russian default of 1998, were all seen as instances where significant contagion did occur. However, with the exception of the Russian default – which affected all emerging and developing regions, as well as the developed world to a surprising extent (Davis, 1999) - the resultant contagion was restricted to the same region. Consequently, a portfolio diversified across all emerging and developing regions would not have suffered simultaneous problems to the extent described above. On the other hand, more recent events, such as the Brazilian devaluation of 1999, Turkey’s devaluation in early 2001 and the problems starting in Argentina towards the end of 2001, have been associated with far less contagion, and have not become an emerging market-wide phenomenon.

Kaminsky et al (op. cit) suggest that a crisis that spreads beyond regional boundaries requires an investment boom, or bubble, to precede it. In this way, actors beyond the region become involved in events there, and so the crisis may spread – via common creditors to some extent – to other emerging, and even developing regions. The current environment is certainly not one of boom with regard to capital flows to emerging and developing economies. Furthermore, it seems unlikely that such circumstances are likely to reoccur in the foreseeable future, ensuring that the preconditions required for system-wide contagion are not in place, and the benefits of widespread diversification will remain a reality.

Kaminsky and Reinhart (2002) also emphasise this point. Their research suggests that financial turmoil in the ‘periphery’ (developing countries) only has systemic implications, such as contagion beyond the immediate region, when asset markets in one of the financial centres (developed world) is affected. “Thus, financial centers serve a key role in propagating financial turmoil. When financial centers remain safe, problems in an emerging market stop at the region’s border”. (p.3)
IV. Conclusion

The expressed purpose of the proposed new Basel Capital Accord is to better align regulatory capital with actual risk. This process, it is argued, will put bank lending on a sounder regulatory footing and remove the many distortions that have come to be recognised in the existing accord. We have argued that the current proposals run the risk of causing an increase in cost and/or reduction in quantity of bank lending to developing countries, as a consequence of the sharp increase in capital requirements for lending to lower rated borrowers. The response to this argument is that any changes in capital requirements are justified on the basis that, whilst the capital associated with lower (higher) rated borrowers is to rise (fall) significantly, relative to the existing situation, this merely reflects the more accurate measurement of risk.

However, as we have demonstrated in this paper, the failure of the proposals to date to take account of the benefits of international diversification suggests that, in this instance at least, risk is not been accurately measured. That is, by excluding the possibility that banks’ capital requirements should take account of portfolio and diversification effects, the proposals effectively impose an inaccurate measure of actual risk, at the portfolio level. At present, the most sophisticated banks often do take account of the benefits of diversification in their international lending decisions. The fact that the proposals under Basel II will not allow these diversification benefits to be taken into account, suggests that the regulatory capital associated with lending to developing countries will be higher than that which the banks would – and currently are – choosing to put aside on the basis of their own models.

The Basel Committee has already made a number of modifications to the original proposals of January 2001 (CP2). The most significant being the modifications to the IRB formula to take account of variable asset correlation as related to PD, and those relating to SMEs. Following the release of CP2 there was widespread concern that lending to SMEs would be adversely affected by a large increase in the capital requirements associated with such lending. After intensive lobbying the Basel Committee has reconsidered the issue. The general changes to the IRB formula with respect to corporate lending – wherein the curve has been significantly flattened – will obviously be of benefit to SMEs. However, the Basel Committee has gone further. July 2002 saw the release of a document by the Basel Committee, which highlighted major areas where agreement had been reached. Of these, it was agreed that the treatment of SMEs should be as follows:

In recognition of the different risks associated with SME borrowers, under the IRB approach for corporate credits, banks will be permitted to separately distinguish loans to SME borrowers (defined as those with less than Euro 50 mn in annual sales) from those to larger firms. Under the proposed treatment, exposures to SMEs will be able to receive a lower capital requirement than exposures to larger firms. The reduction in the required amount of capital will be as high as twenty percent, depending on the size of the borrower, and should result in an average reduction of approximately ten percent across the entire set of SME borrowers in the IRB framework for corporate loans.\textsuperscript{11}

Thus, in the case of SME and corporate lending, the Basel Committee has recognised the impact that differential asset correlation can have on portfolio level risk. Our

\textsuperscript{11} Basel Committee reaches agreement on New Capital Accord issues. http://www.bis.org/press/p020710.htm
results strongly suggest that a similar modification is justified with respect to internationally diversified lending.

The specific manner that the Basel Committee might want to incorporate these findings is, of course, best left to them. Given the experience and expertise at their disposal we would not at this stage want to offer suggestions as to the means by which these modifications might be made. However, given the changes already made to the IRB formula with respect to corporates and SMEs, as well as the fact that the changes we propose would seem to have at least as solid an empirical basis, there are no theoretical, empirical or practical reasons why changes should not be made to incorporate the benefits of international diversification. We therefore urge the Basel Committee to incorporate these findings in the final consultative paper, due for release in Spring 2003, and would be happy to collaborate with the Committee in this important work, if it was considered useful.

References:


Figure 1. C.D.F Test for Correlations on Syndicated Loan Spreads (1993-2002)

Figure 2. C.D.F Test for Correlations on Banks’ Return on Capital (1988-2001)

Figure 3. C.D.F Test for Correlations on Banks’ Return on Assets (1988-2001)

Figure 4. C.D.F Test for Correlations on GDP Growth (1985-2000)

Figure 5. C.D.F Test for Correlations on Real Short-Term Interest Rates (1985-2000)

Figure 6. C.D.F Tests for Correlations on Stock Exchange Movements (IFC I-COMP: 1990-2002)

Figure 7. C.D.F Tests for Correlations on Stock Exchange Movements (IFC G-COMP: 1990-2002)

Figure 8. C.D.F Tests for Correlations on Bond Market Movements (GBI-EMBI+: 1991-2002)
Appendix: Computation of Unexpected Losses

Considering that the quality of the credit portfolio of any bank can change at any time in the future, there is a need to make frequent calculations of the expected losses that a bank could suffer under different risk situations. Given the constant changes in portfolio quality, it is unlikely that the computed expected losses will be the same for different periods. The difference between expected losses computed at different periods, (due to changing credit quality), is the cause of potential losses to the bank, that could erode their capital in extreme situations. These losses are called “Unexpected Losses” and their estimation represents the issue to be addressed in this appendix.

Unexpected Losses arise because of joint credit quality changes among the credits that conform a credit portfolio. In order to model such joint credit quality changes, we adopted a portfolio approach.

The adoption of the portfolio approach (See Markowitz (1959)) has been amply documented and adopted in diverse finance applications. Under this theory, investors formulate their investment portfolio, taking care of the optimal risk-return relationship that a given portfolio has. With this in mind, credit risk modellers have already developed risk management techniques aimed to take account of the portfolio diversification effect. Although such approaches might be subject to improvements, we do believe that portfolio diversification could and should be an integral part of credit risk valuation for regulation purposes. As we have argued in the main body of paper and in previous work, we believe that negative economic outcomes will be provoked by the fact that the proposed regulation framework only punishes high-risk taking but does not provide incentives for portfolio diversification.

In this appendix we present a modification to the “Credit-Metrics” methodology that has been used to simulate credit unexpected losses of analysed portfolios. The Credit-Metrics model has been described as: “A full portfolio view addressing credit event correlations which identify the costs of over concentration and benefits of diversification”. (See J.P. Morgan (1997)). The objective of this appendix is to present the modifications that were made to the Credit-Metrics approach in order to make possible its implementation. For detailed exposition refer to the Credit-Metrics technical document. We refer to the modification to the “Credit-Metrics” methodology as Full Credit Risk model: FCRM.

Full Credit Risk Model:

Empirical studies that show that credit defaults are correlated have been widely presented, here we present evidence that credit risk can also be diversified. In order to calculate portfolio diversification, it would be necessary to know the probability that each of the credits making up a portfolio migrate jointly from their current rating (credit quality) to each of the possible ratings. For this, we would

\[\text{If it is not our intention at this point to analyse the possible improvements to each methodology.}\]

\[\text{The choice of this model was made in terms of simplicity for modelling and availability of data. It is not our intention to favour any specific credit risk modelling technique.}\]
require to know a number of tables of joint probabilities equivalent to the number of pairs of credits making up a portfolio. This objective is unattainable given the lack of reliable data, the amount and complexity of it.

The CreditMetrics’ approach makes use of two main elements:

I. The Merton Approach to model Credit Quality Changes.
II. An indirect approach to model Correlations among the credits that make up a credit portfolio.

Finally, once a correlations matrix among the creditors making up the credit portfolio is built, this methodology simulates the unexpected losses for the portfolio.

I. The Merton Approach to model Credit Quality Changes.

The Merton approach assumes that equity can be viewed as a call option on the firms’ assets with a strike price equal to the book value of the firm’s debts (See Merton (1974)). The intuition behind this assumption is that given the limited liability feature of equity, equity holders have the right but not the obligation to payoff debt-holders and take over the remaining assets of the firm. This approach implies that the credit quality (rating) of a given creditor is related to the difference between the market value of its assets and its debt.

Under this approach, the change in the value of the assets of a given company is related to the change in its rating. Therefore, the distribution of the company’s assets returns can be used to calculate the distribution of its rating change's probabilities. For the generalisation of this model, it is necessary to include, in addition to the default state, different credit quality states.

Figure A.1: The Distribution of Assets’ Returns.

The transition matrix is the variable that summarises the migration probabilities from one credit quality to any other. Having the transition probabilities between different credit qualities and considering the Merton Approach, it is possible to derive the market value of assets that represent the cut-off values between different credit qualities, as shown in Figure 1. These cut-off values fulfil the condition that if the change in the market value of the asset (r) is sufficiently negative, (i.e. smaller than ZE ), then the credit falls into default; if ZE < r < ZD, the credit is rated D, and so on.
Taking into consideration the empirical transition matrix, it is possible to estimate the probability of these changes happening as follows: (for a credit initially rated as X).

\[
\begin{align*}
\text{Prob}(E|X) &= \text{Prob}(r < ZE) = \phi(ZE) \\
\text{Prob}(D|X) &= \text{Prob}(ZE < r < ZD) = \phi(ZD) - \phi(ZE) \\
\text{Prob}(C|X) &= \text{Prob}(ZD < r < ZC) = \phi(ZC) - \phi(ZD) \\
\text{Prob}(B|X) &= \text{Prob}(ZC < r < ZB) = \phi(ZB) - \phi(ZC) \\
\text{Prob}(A|X) &= \text{Prob}(ZB < r < ZA) = 1 - \phi(ZB)
\end{align*}
\]

Where:
\(R\): Is the implied market value of assets.
\(\phi\): Is the cdf for the Normal distribution

From this point of view, the correlation matrix of changes of credit quality between creditors can be computed by developing an explanatory model of the changes in the value of the assets of the debtors.

This approach presents several practical problems for implementation, the most important being the handling of very large correlation matrices. Additionally, it is not possible to obtain the changes in the market value of assets for each particular debtor, since it would be necessary to have specific information about the internal financial structure of each debtor. These two disadvantages make it impossible to implement an ideal correlation matrix, for these reasons we will adopt an indirect (but more manageable) method to introduce the portfolio diversification effect.

II. An indirect approach to model Correlations among the credits that make up a credit portfolio.

Following the Merton’s approach, J.P. Morgan makes an a-priori distinction of the factors that determine the changes in the value of the assets of the debtors. This distinction comes from two basic components: the market component and the idiosyncratic component. By definition, the idiosyncratic component does not correlate with anything, since it refers to those factors unique for the debtor. But the market component has with it, all the elements that allow the portfolio diversification.

\[
r_{\text{total}} = W_M r_{\text{market}} + W_I r_{\text{idiosyncratic}}
\]

Where:
\(W_M\): Percentage of returns explained by the market component\(^14\).
\(r_{\text{market}}\): Market component of returns.
\(W_I\): Percentage of returns explained by the idiosyncratic component\(^15\).
\(r_{\text{idiosyncratic}}\): Idiosyncratic component of returns.

\(^14\) In the CreditMetrics technical document, it is explained how these weights can be calculated. After empirical implementations, it is proved that an acceptable value of \(W_M = 70\%\). For our exercise, we assume this value.

\(^15\) The idiosyncratic component weight is obtained with the following equation:

\[
W_I = \sqrt{1 - w_M^2}
\]

The objective of this equation is to be consistent with the change in the market value of the assets’ standardized returns.
Conversely, the market component of Returns is defined as:

\[ r_{\text{Market}} = H_A \cdot r_{\text{GDP Country}} + (1-H_A) \cdot r_{\text{GDP Economic Activity}} \quad (2) \]

Where:
- \( H_A \): Percentage of market component explained by the GDP of the debtor’s country. The Herfindahl Index computes this parameter.
- \( r_{\text{GDP Country}} \): Return on the GDP of the debtor’s country.
- \((1-H_A)\): Percentage of market component explained by the GDP of the debtor’s country.
- \( r_{\text{GDP Economic Activity}} \): Return on the GDP of the debtor’s Economic Activity.

The market component of returns is divided between economic activity and geographical area. Which is more relevant for a debtor? Is it his economic activity or the country where his business takes place? The percentage of participation of these market factors in the debtor’s systemic risk is exogenous to the model. Therefore a methodology was designed to solve this problem in the most objective way possible (See Segoviano (1998)).

This methodology was based in the fact that the greater the variety of economic activities in a country, the lesser the effect (on the value of assets of a debtor in that country) of a sudden change in the country’s production. Within this framework it is possible to infer that in those countries in which there are a few economic activities (and therefore there is a high economic activity concentration), the most important factor for the debtor’s assets value will be his geographic location. The intuition behind this reasoning is the fact that if the country is affected by an economic shock, it is very likely that debtor will experience a decrease in the value of his assets, since it is highly probable that the debtor will belong to the economic activities that have been affected.

Following this reasoning, we computed a “Herfindahl” index with the following formula for each group of countries:

\[ H_A = \sum_{i=1}^{n} \left( \frac{X_{Ai}}{\sum_{j=1}^{n} X_{Aj}} \right)^2 \tag{3} \]

Where:
- \( X_{Ai} \): is the amount of participation of the i economic activity in country group A.  

Once considered all the elements that compose the market component of assets’ returns the next step is to calculate the correlations between the debtors making up a credit portfolio.

Given a pair of debtors X and Y, working in B and V industrial activities; located in A and E country groups and with returns expressed in the following way:

\[ 16 \]

The higher the Herfindhal index for a given country group, the less economic activity diversified. Then, the percentage of the market component explained by the GDP of the debtor’s country takes more importance.
The problem of estimating the correlations among each couple of creditors in the portfolio is summarised in the following way:

\[ \rho_{XY} = w_{MX}^2 H_A r_A + w_{MY}^2 H_E r_E + w_{MX} w_{MY} (1 - H_A) \rho_{BV} \]  \hspace{1cm} (4)

Where:

\( \rho_{AE} \): is the correlation between different country groups\(^{18} \).

\( \rho_{BV} \): is the correlation between different economic activities\(^{19} \).

This equation is computed for each pair of debtors making up the portfolio. The results of computing this equation are compiled in a \((n \times n)\) square matrix, where \(n\) is the number of creditors in the portfolio. This matrix is named the correlation matrix between creditors and is unique for each portfolio. This matrix is an extremely important variable for the simulation of unexpected losses, since it incorporates the necessary elements to quantify the concentration/diversification of the portfolio.

With these elements, we show in the following section how quality scenarios for the portfolio are simulated. From these quality scenarios, the loss distribution is built from which it is possible to obtain the unexpected losses.

### III Simulation of quality scenarios for the credit portfolio

Combining the transition matrix with the correlation matrix between creditors, we simulate quality scenarios from which the loss distribution for the credit portfolio is obtained.

As explained above, the transition matrix indicates the probabilities of quality changes that a creditor with a given rating might experience. Additionally, the correlations of quality changes between creditors is involved. Creditors with similar characteristics will tend to migrate jointly to different credit qualities when hit by economic shocks. Creditors with different characteristics will tend to migrate dis-jointly to different credit qualities when hit by economic shocks. This implies that credit portfolios concentrated in credits with similar characteristics will tend to have higher unexpected losses since they will not be diversifying the possible economic risks.

We programmed an algorithm to compute 10,000 possible quality scenarios for each of the \((n \times n)\) couples of creditors that make up the portfolio. Each quality scenario

\(^{17}\) Since the correlations between idiosyncratic components and geographical components, idiosyncratic components and economic activity components as well as between economic activities and geographical components are assumed to be zero.

\(^{18}\) These correlations were computed between the spreads of syndicated loans for each country group. We considered that such spreads represent the riskiness of the financial system in each country group.

\(^{19}\) These correlations were computed between indexes for each of the economic activities considered in the exercise. Each economic activity index was built with the economic activity component of the GDP of a representative country for each country group in the sample.
shows a change in the market value of the assets of the creditors in the portfolio. This process is repeated 10,000 times. The quality changes of the members of the portfolio allow generating an amount of losses or profits that conform the loss distribution of the portfolio.

In order to generate these scenarios, the following process is computed:

1. Generation of random uniform numbers.
2. Transformation of this random numbers into normal standard random numbers.
3. Transformation of the normal standard random numbers into normal multivariated random numbers with variance equal to the correlation matrix between creditors.

Since it was assumed that the process that generates changes in the assets follows a normal distribution, we use normal random multivariated distribution to generate joint quality migrations, where credits with high correlation will tend to migrate jointly.

IV Unexpected Loss

Once the credit portfolio quality scenarios have been simulated, it is possible to compute the losses/gains that come from the difference between initial and final credit qualities. The losses/gains obtained from the simulation process are used to build a histogram. This histogram summarizes the loss distribution of the credit portfolio.

Figure A2: Credit Portfolio Loss Distribution

Simulated unexpected losses must be ordered to generate the loss distribution.

From this distribution a Value at Risk (VaR) is defined from which we obtain the amount of unexpected losses from the portfolio. The unexpected losses divided by the total amount of the portfolio represent the percentage that with a given probability (defined by the chosen percentile) could be lost in an extreme event. Then, capital requirements should be such that they can cover such losses.